Provider and Hospital Characteristics Associated With Geographic Variation in the Evaluation and Management of Elderly Patients With Heart Failure

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Background: Rates of guideline-based care for elderly patients with heart failure vary by state, and overall are not optimal. Identifying factors associated with the lack of uniformly high-quality health care might aid efforts to improve care. We therefore sought to determine the extent to which provider and hospital characteristics contribute to small-area geographic variation in heart failure care after controlling for patient factors.

Methods: We studied 30,228 Medicare patients who were older than 65 years and hospitalized with heart failure. We mapped rates for two quality measures—documentation of left ventricular ejection fraction and appropriate prescription of angiotensin-converting enzyme inhibitors—across the United States, using a Bayesian technique that smooths rates and enhances assessment for significant patterns of small-area variation. We used nonlinear hierarchical models to assess for associations between the quality indicators and provider and hospital characteristics independent of patient characteristics.

Results: Smoothed, unadjusted rates of left ventricular ejection fraction documentation ranged from 30.1% to 67.2% and of angiotensin-converting enzyme inhibitor prescription from 55.8% to 87.1% among hospital referral regions; regional patterns were apparent. After patient factors were controlled for, care at hospitals without a medical school affiliation, without invasive cardiac capabilities, or in a rural location, as well as not having a cardiologist as an attending physician, was significantly associated with lower rates of left ventricular ejection fraction documentation. Hospitalization at a nonteaching facility was significantly associated with failure to prescribe angiotensin-converting enzyme inhibitors.

Conclusion: Characteristics of providers and hospitals explain in part the geographic variation in guideline-based care for elderly patients with heart failure.

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ure care. We assessed the degree of small-area geographic variation for these indicators by mapping rates with a Bayesian technique to facilitate the recognition of geographic patterns by hospital referral region (HRR). Small areas such as HRRs have been judged to be the most valid units for characterizing geographic variation. We also looked for associations between indicator rates and hospital and physician characteristics after controlling for patient characteristics.

QUALITY OF HEART FAILURE CARE

Our principal outcomes were 2 indicators of the quality of heart failure care. The first indicator assessed whether patients had documentation in the medical record of left ventricular EF assessment performed either before or during the index admission. The second indicator assessed whether patients with left ventricular systolic dysfunction (EF < 40% or the narrative equivalent) were prescribed an ACE inhibitor at the time of discharge or had a documented reason why ACE inhibitors were not used. Both measures are widely supported in evidence-based guidelines for heart failure care.

PATIENT SAMPLE

We used data from the Centers for Medicare & Medicaid Services’ National Heart Care Project. The heart failure component of the project, known as the National Heart Failure Project, was designed to improve the quality of care for Medicare beneficiaries with heart failure. As part of the project, data were abstracted from the hospital charts of patients discharged between April 1998 and March 1999 with a principal diagnosis of heart failure (International Classification of Diseases, Ninth Revision codes 402.01, 402.11, 402.91, 404.01, 404.11, 404.91, and 428.x); code 428.x has a specificity of more than 95% for the diagnosis of heart failure. Patients were excluded if they left the hospital against medical advice, were transferred to another acute care facility, were receiving long-term hemodialysis, or had no Social Security number recorded. One discharge was chosen randomly for inclusion if there were multiple discharges within a state for the same patient.

A random sample of 800 or 900 charts per state was drawn after all available charts from that state were stratified by patient age, race, sex, and hospital. All charts were used when fewer than 800 charts were available from a state. Trained medical chart abstractors at 2 central locations abstracted the data.

DATA SOURCES AND DEFINITIONS

Patient clinical data were obtained from chart review. Comorbidity conditions were recorded as present if the diagnosis was present in the admission, consultation, or progress notes. We defined patient socioeconomic status with 2 measures. The first was dual eligibility, indicating that the patient was eligible for both Medicare and Medicaid, and the second was the ZQ score for the patient’s ZIP code. The ZQ score (Claritas Inc, San Diego, Calif), which is derived from census data, is calculated from household income, educational attainment, occupation scores, and home values aggregated by ZIP code and ranges from 1 (lowest socioeconomic status) to 100 (highest possible socioeconomic status). Patients were classified as Hispanic based on information in the medical record and on use of a surname that was consistent with the Hispanic surname from the patient’s Hispanic surname from the patient’s Social Security number or name on social security document number. Patient-level data elements included demographic data, heart failure symptoms, comorbid conditions, physical examination findings, and laboratory data, including electrocardiographic and chest radiograph findings. Variables associated in univariate analysis with a P value of less than .05 were candidates for the multivariable models; we chose the relatively stringent criterion of P < .05 because of the large number of observations in our sample. We included variables identified as clinically important from prior studies without subjecting them to univariate analysis. These variables included attending physician specialty, teaching hospital status, and the presence of cardiac catheterization facilities. We estimated the parameters for the covariates in models that included a ran-

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STATISTICAL METHODS

Variation in rates among small areas is the result of both random variation within areas around the true rate for that area and systematic variation in true rates across areas. The literature on small-area variation has produced a number of measures that describe this variation across areas but has not produced clear consensus on the best measure. Recently, empirical Bayesian techniques, originally developed to solve image reconstruction problems, have been applied to problems of spatial epidemiology. This technique revises the estimate of the dependent variable for a given region using the mean value of the estimates for the adjacent regions. The empirical Bayesian techniques has advantages over traditional maximum likelihood techniques in that it includes more reliable accounting for variation in areas of low incidence or of relatively high or low population and it produces more easily interpretable maps. We used a nonlinear mixed-effects model with HRR designated as the random effect to produce the empirical Bayesian estimates of the rates by HRR. The parallel between mixed-effects models, also known as hierarchical models, and Bayesian statistics is the result of the correspondence between the random coefficients in the mixed-effects models and the random parameters essential to Bayesian statistics.

The null model includes random effects but no covariates. From this, we estimated the unadjusted rates by HRR. The smoothing is the result of shrinkage toward the mean, since the posterior means are a weighted combination of the population estimates and the observed data, where the weights are a function of the variance of the random effects (in this case, a random intercept for each HRR) and the residual error. The emphasis is on the population component for HRRs with few observations and shifts toward the observed component as the number of observations increases.

We added covariates to the null hierarchical models to describe the relationships between rates of EF documentation or ACE inhibitor prescription and characteristics of patients, hospitals, and providers. These covariates were chosen from the 192 patient-level data elements in the National Heart Failure database as well as from the provider variables available in the American Medical Association Physician Masterfile and the hospital variables from the American Hospital Association annual survey as noted above. Patient-level data elements included demographic data, heart failure symptoms, comorbid conditions, physical examination findings, and laboratory data, including electrocardiographic and chest radiograph findings. Variables associated in univariate analysis with a P value of less than .05 were candidates for the multivariable models; we chose the relatively stringent criterion of P < .05 because of the large number of observations in our sample. We included variables identified as clinically important from prior studies without subjecting them to univariate analysis. These variables included attending physician specialty, teaching hospital status, and the presence of cardiac catheterization facilities. We estimated the parameters for the covariates in models that included a ran-

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RESULTS

There were 37,423 unique patients in the analysis sample. After the inclusion and exclusion criteria were applied, 30,228 patients aged 65 years or older were eligible for the ACE documentation indicator, and 8,475 patients aged 65 years or older were eligible for the ACE inhibitor prescription indicator. Exclusions are shown in Figure 1. The initial subset of 30,817 had an age distribution typical for Medicare patients (29.6%, 65-74 years; 43.5%, 75-84 years; and 26.9%, ≥85 years) and was 59.7% female. The patient variables associated with documentation of EF before or during hospitalization were female sex, older age, lower income (as assessed by Medicaid eligibility), higher serum creatinine levels, diabetes, chronic obstructive lung disease, history of stroke or dementia, and absence of history of hypertension. The attending physician characteristic of cardiology specialty was associated with more frequent documentation of EF. The hospital characteristics associated with lower rates of documentation of EF were smaller size, absence of catheterization facilities, nonteaching hospital status, and rural location. The univariate relationships for all candidate variables are not shown because of the large number of variables investigated.

Odds ratios for the association of patient, physician, and hospital characteristics with rates of ACE inhibitor prescription are shown in Table 2. The patient variables associated with failure to prescribe ACE inhibitors were older age, nonblack race, higher serum creatinine levels, and chronic obstructive lung disease. No physician characteristics were associated with failure to prescribe ACE inhibitors, but the hospital characteristic of nonteaching hospital designation was associated. Again, univariate relationships for all candidate variables are not shown because of the large number of variables investigated.

Table 3 shows odds ratios for US Census regions (East, Midwest, South, and West) with East as the reference variable. This component of the analysis quantifies the regional patterns evident in Figures 2 and 3, with significantly higher rates of EF determination in the Northeast compared with the rest of the country, and signifi-
significantly lower rates of ACE inhibitor prescription in the South.

Summary statistics for the models are given in Table 4. The statistically significant Hosmer-Lemeshow statistic demonstrates that the model is relatively poorly calibrated and should not be used to predict quality indicators from patient and provider characteristics, a purpose for which it was not intended.

The current study demonstrates small-area variation in rates of EF documentation and ACE inhibitor prescription for hospitalized elderly patients with heart failure in the United States. Variation in rates smoothed with a Bayesian technique is strong evidence that this variation is not the result of chance. Rates of these quality indicators were generally highest in the northeastern states and generally lowest in the Plains and southern states. Rates were lower and variation was greater for EF documentation than for ACE inhibitor prescription.

After patient characteristics were adjusted for, both provider and hospital characteristics remained significantly associated with EF documentation but not with

### Table 1. Multivariable Relationship Between Rates of Ejection Fraction Determination and Patient, Physician, and Hospital Characteristics

<table>
<thead>
<tr>
<th>OR (95% CI)</th>
<th>Female</th>
<th>0.86 (0.82-0.91)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age, y</td>
<td>75-84</td>
<td>0.88 (0.83-0.94)</td>
</tr>
<tr>
<td></td>
<td>&gt;85</td>
<td>0.62 (0.58-0.68)</td>
</tr>
<tr>
<td>Dually eligible</td>
<td>0.78 (0.73-0.83)</td>
<td></td>
</tr>
<tr>
<td>Serum creatinine*</td>
<td>0.87 (0.84-0.91)</td>
<td></td>
</tr>
<tr>
<td>History of hypertension</td>
<td>1.29 (1.22-1.36)</td>
<td></td>
</tr>
<tr>
<td>COPD</td>
<td>0.81 (0.76-0.85)</td>
<td></td>
</tr>
<tr>
<td>History of stroke</td>
<td>0.84 (0.79-0.90)</td>
<td></td>
</tr>
<tr>
<td>Diabetes</td>
<td>0.87 (0.82-0.92)</td>
<td></td>
</tr>
<tr>
<td>Dementia</td>
<td>0.71 (0.66-0.78)</td>
<td></td>
</tr>
<tr>
<td>Cardiac catheterization laboratory present</td>
<td>1.19 (1.10-1.29)</td>
<td></td>
</tr>
<tr>
<td>Teaching hospital</td>
<td>1.12 (1.03-1.22)</td>
<td></td>
</tr>
<tr>
<td>Rural</td>
<td>0.58 (0.53-0.65)</td>
<td></td>
</tr>
<tr>
<td>Hospital size†</td>
<td>1.31 (1.24-1.39)</td>
<td></td>
</tr>
<tr>
<td>Cardiologist as attending physician</td>
<td>1.17 (1.10-1.26)</td>
<td></td>
</tr>
<tr>
<td>Attending physician specialty unknown</td>
<td>1.17 (1.07-1.28)</td>
<td></td>
</tr>
</tbody>
</table>

### Table 2. Multivariable Relationship Between Rates of ACE Inhibitor Prescription and Patient, Physician, and Hospital Characteristics

<table>
<thead>
<tr>
<th>OR (95% CI)</th>
<th>Black race</th>
<th>1.35 (1.12-1.62)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age, y</td>
<td>75-84</td>
<td>0.86 (0.75-0.97)</td>
</tr>
<tr>
<td></td>
<td>&gt;85</td>
<td>0.83 (0.71-0.96)</td>
</tr>
<tr>
<td>Serum creatinine*</td>
<td>0.56 (0.52-0.60)</td>
<td></td>
</tr>
<tr>
<td>COPD</td>
<td>0.79 (0.71-0.89)</td>
<td></td>
</tr>
<tr>
<td>Teaching hospital</td>
<td>1.20 (1.07-1.35)</td>
<td></td>
</tr>
</tbody>
</table>

### Table 3. Odds Ratios (OR) for Ejection Fraction (EF) Documentation and ACE Inhibitor Prescription by US Census Regions

<table>
<thead>
<tr>
<th>OR (95% CI)</th>
<th>EF Documented</th>
<th>ACE Inhibitor Prescribed</th>
</tr>
</thead>
<tbody>
<tr>
<td>East</td>
<td>71.8</td>
<td>83.8 Referent</td>
</tr>
<tr>
<td>Midwest</td>
<td>64.1</td>
<td>80.4 0.70 (0.61-0.80)</td>
</tr>
<tr>
<td>South</td>
<td>64.9</td>
<td>76.5 0.72 (0.64-0.82)</td>
</tr>
<tr>
<td>West</td>
<td>62.3</td>
<td>82.7 0.64 (0.55-0.77)</td>
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</tbody>
</table>

### Table 4. Model Summary Statistics

<table>
<thead>
<tr>
<th>R², %</th>
<th>c Statistic†</th>
<th>H-L χ²‡ (P Value)</th>
</tr>
</thead>
<tbody>
<tr>
<td>EF documentation</td>
<td>10.5</td>
<td>0.69</td>
</tr>
<tr>
<td>ACE inhibitor prescription</td>
<td>6.4</td>
<td>0.63</td>
</tr>
</tbody>
</table>

Abbreviations: CI, confidence interval; COPD, chronic obstructive pulmonary disease; OR, odds ratio.

*Odds ratio is for each 1-mg/dL increment in serum creatinine.

†Logarithm of number of beds.

The current study demonstrates small-area variation in rates of EF documentation and ACE inhibitor prescription for hospitalized elderly patients with heart failure in the United States. Variation in rates smoothed with a Bayesian technique is strong evidence that this variation is not the result of chance. Rates of these quality indicators were generally highest in the northeastern states and generally lowest in the Plains and southern states. Rates were lower and variation was greater for EF documentation than for ACE inhibitor prescription.

After patient characteristics were adjusted for, both provider and hospital characteristics remained significantly associated with EF documentation but not with

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ACE inhibitor prescription. Contributions of provider and hospital characteristics were quantitatively similar. The effect of adding hospital characteristics was more pronounced in the model for EF measured than in the model for ACE inhibitors, while the addition of patient characteristics caused the single largest reduction in unexplained variation in the ACE inhibitors model.

Obviously, EF documentation and ACE inhibitor prescription are not independent quality indicators; the selection of patients with heart failure for ACE inhibitor therapy depends on first determining the EF. Thus, in areas where the EF is less likely to be documented, opportunities are missed for prescribing proper therapy.

Several of these findings bear further comment. The finding of chronic obstructive pulmonary disease as a predictor of both failure to document EF and failure to prescribe an ACE inhibitor is likely the result of diagnostic overlap between this diagnosis and heart failure. The finding of black race as a predictor of ACE inhibitor prescription is surprising in light of a perception of a lack of efficacy for these agents in African American patients. Although this finding may be the result of African American patients being more likely to have hypertension requiring treatment with multiple agents, further investigation is needed. The finding of lower rates of EF documentation in rural hospitals and in those without invasive cardiac services may imply that echocardiography is unavailable in these settings, although we attempted to compensate for this by excluding from the denominator cases with documentation of a plan to measure EF after discharge.

**COMPARISON WITH PREVIOUS LITERATURE**

Geographic variation in patterns of care for heart failure has not previously been described, to our knowledge. There are parallels in care for ischemic heart disease, however, between the findings in the current study and the findings in studies of geographic variation.

In studies of Medicare patients in the Cooperative Cardiovascular Project, quality of care varied across the country. Higher quality of care after myocardial infarction in Medicare patients has been associated with male sex and higher income, younger age, teaching hospitals, and care by a cardiologist. Rates have been highest in New England.

There is one important way, however, in which our findings differ from those of services for ischemic heart disease. Studies of the relationship between quality of care for ischemic heart disease and the availability of invasive cardiac services have had ambiguous results, and much of the literature has implied that availability of subspecialty services is associated with overuse of services. In the current study, however, facilities with catheterization laboratories and attending physicians whose specialty is cardiology were more likely to provide guideline-based care.

**LIMITATIONS OF CURRENT STUDY**

Our study has several potential limitations. We did not seek to confirm the diagnosis of heart failure. Given that both overdiagnosis and underdiagnosis of heart failure have been documented in hospital charts, restricting the diagnosis to patients with clear evidence of heart failure would necessarily trade lower sensitivity for higher specificity and exclude many patients with heart failure. Ascertainment of patient socioeconomic status is difficult; in our study, the patient-specific variable of dual eligibility for Medicare and Medicaid performed better than a ZIP code level variable accounting for income, occupation, education, and home values. Our study was restricted to hospitalized patients. This group, however, arguably deserves the most attention with regard to quality improvement. We report odds ratios and not relative risks; therefore, the magnitudes of the demonstrated associations should be interpreted with caution.

Finally, important contributors to the variation in quality of care including patient factors were unavoidably omitted from the analysis. This view is supported by the relatively low values for the c statistic for both models; taken in toto, patient, hospital, and provider characteristics explain geographic variation modestly. Because quality of care is affected by organizational factors such as the use of data feedback, the influence of opinion leaders, organizational stability, and the presence of administrative support—factors that cannot be measured using nationwide data—unexplained variation should be expected.

**CONCLUSIONS AND IMPLICATIONS FOR POLICY**

Geographic variation in the provision of high-quality heart failure care cannot be completely explained by patient characteristics. Characteristics of the health care system such as availability of cardiology specialty services play important roles. Unmeasured, and for the purposes of this study immeasurable, factors such as the presence of a “culture of quality” within hospitals likely also play important roles. Further understanding of these issues will provide useful context for public reporting of hospital performance. Future research should focus on characterizing these unmeasured factors further and on identifying innovative systems to make the characteristics of the best-performing hospitals available to all facilities.

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